

STAT 24620=FINM 34700, STAT 32950
Multivariate Data Analysis
Lecture 2: Principal Component Analysis: Foundations

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Outline

- 1 Bridge from Lecture 1 + Motivation
- 2 PCA Optimization Formulation
- 3 Variance, Geometry, and SVD
- 4 Practical Decisions + Financial Example
- 5 Wrap-up

Bridge from Lecture 1: projection language and notation

- Observed data matrix: $\mathbf{X} \in \mathbb{R}^{n \times p}$.
- Centered matrix: \mathbf{X}_c (each column mean removed).
- A projection direction is denoted by $\mathbf{v} \in \mathbb{R}^p$ with $\|\mathbf{v}\|_2 = 1$.

Projected score (observation i):

$$z_i = \mathbf{x}_{c,i}^\top \mathbf{v} = (\mathbf{x}_i - \bar{\mathbf{x}})^\top \mathbf{v}.$$

Key question for today:

Among infinitely many unit directions \mathbf{v} , how to choose the most informative ones?

Population PCA vs Sample PCA

Population version (theoretical target)

- Random vector $X \in \mathbb{R}^p$ with mean μ and covariance Σ .
- Population PCs aims to maximize the true variance $\text{Var}(X^\top \mathbf{v}) = \mathbf{v}^\top \Sigma \mathbf{v}$.

Sample version (what we actually compute)

- Observations $\mathbf{x}_1, \dots, \mathbf{x}_n$ and centered matrix \mathbf{X}_c .
- Sample covariance:

$$\mathbf{S} = \frac{1}{n-1} \mathbf{X}_c^\top \mathbf{X}_c.$$

- Sample PCs maximizes the sample variance $\widehat{\text{Var}}(X^\top \mathbf{v}) = \mathbf{v}^\top \mathbf{S} \mathbf{v}$.

Key idea: sample PCA is an estimate of population PCA.

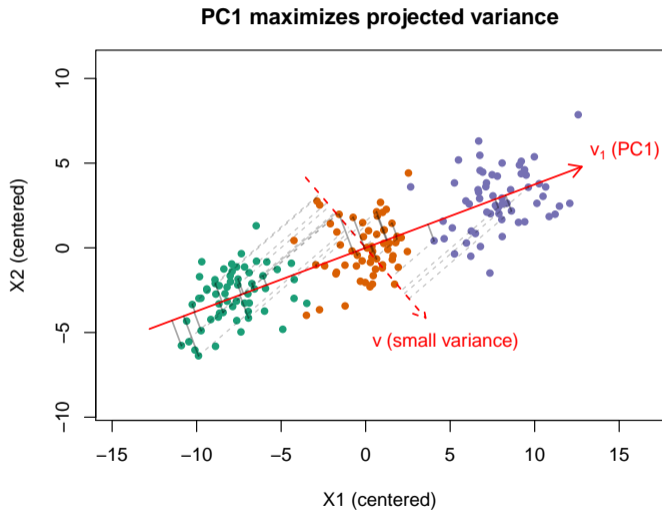
Why maximize variance?

- A direction with very small projected variance makes the data look almost flat.
- A direction with large projected variance reveals major heterogeneity in the data.
- If we keep only a few components, we want them to retain as much information as possible.

Compression perspective:

- PCA chooses directions that minimize information loss under squared reconstruction error.
- Thus, “maximum variance” and “best low-rank approximation” are two sides of the same coin.

Data cloud and first PC direction



Sample PCA optimization for the first component

For a unit direction \mathbf{v} , sample projected variance is

$$\widehat{\text{Var}}(X^T \mathbf{v}) = \mathbf{v}^T \mathbf{S} \mathbf{v}.$$

Hence, the first sample PC solves

$$\max_{\mathbf{v} \in \mathbb{R}^p} \mathbf{v}^T \mathbf{S} \mathbf{v} \quad \text{s.t.} \quad \|\mathbf{v}\|_2^2 = \mathbf{v}^T \mathbf{v} = 1.$$

Interpretation:

- Objective: maximize spread in projected scores.
- Constraint: avoid trivial scaling ($c\mathbf{v}$ would otherwise make objective arbitrarily large).

How do we find the solution?

Lagrangian method (for constrained optimization):

$$\mathcal{L}(\mathbf{v}, \lambda) = \mathbf{v}^\top \mathbf{S} \mathbf{v} - \lambda(\mathbf{v}^\top \mathbf{v} - 1).$$

- We combine objective + constraint using multiplier λ .

first-order condition (FOC):

- At an optimum, derivative of \mathcal{L} w.r.t. \mathbf{v} must be zero:

$$\nabla_{\mathbf{v}} \mathcal{L} = 2\mathbf{S}\mathbf{v} - 2\lambda\mathbf{v} = 0.$$

- So $\mathbf{S}\mathbf{v} = \lambda\mathbf{v}$: optimal directions are eigenvectors of \mathbf{S} .

To maximize $\mathbf{v}^\top \mathbf{S} \mathbf{v} = \lambda \mathbf{v}^\top \mathbf{v} = \lambda$, we take the largest eigenvalue λ_1 .

For PC1: principal components, loadings, and scores

For the first component:

- **Principal component:** the eigenvector \mathbf{v}_1 that satisfies $\mathbf{S}\mathbf{v}_1 = \lambda_1\mathbf{v}_1$.
 - \mathbf{v}_1 may not be unique. (Why?)
- **Loading:** entries of \mathbf{v}_1 ; they show how original variables contribute to PC k .
- **Scores:** projected coordinates of observations on PC direction,

$$z_{i1} = \mathbf{x}_{c,i}^T \mathbf{v}_1.$$

Interpretation: loadings describe variables; scores describe observations.

Why do we need a 2nd component? Why orthogonal?

Why 2nd component?

- PC1 is one-dimensional summary only.
- Residual variation remains after removing PC1 information.
- PC2 captures the strongest remaining variation.

Why orthogonality to PC1?

- To avoid rediscovering the same direction/information.
- Orthogonality ensures non-redundant factors in Euclidean geometry.
- With centered data, orthogonal loading vectors imply uncorrelated PC scores.

Optimization for PC2:

$$\max_{\mathbf{v}} \mathbf{v}^T \mathbf{S} \mathbf{v} \quad \text{s.t.} \quad \|\mathbf{v}\|_2 = 1, \mathbf{v}^T \mathbf{v}_1 = 0.$$

How to find PC2? (short derivation)

let $\mathbf{S} = \mathbf{V}\mathbf{\Lambda}\mathbf{V}^\top$ with

$$\mathbf{V} = [\mathbf{v}_1, \dots, \mathbf{v}_p], \quad \mathbf{V}^\top \mathbf{V} = \mathbf{V}\mathbf{V}^\top = \mathbf{I}_p,$$

$$\mathbf{\Lambda} = \text{diag}(\lambda_1, \dots, \lambda_p), \quad \lambda_1 \geq \dots \geq \lambda_p \geq 0.$$

So \mathbf{V} is the orthogonal matrix whose columns are exactly the eigenvectors $\mathbf{v}_1, \dots, \mathbf{v}_p$.

- Expand any unit \mathbf{v} orthogonal to \mathbf{v}_1 in eigenbasis:

$$\mathbf{v} = \sum_{j=2}^p c_j \mathbf{v}_j, \quad \sum_{j=2}^p c_j^2 = 1.$$

- Then

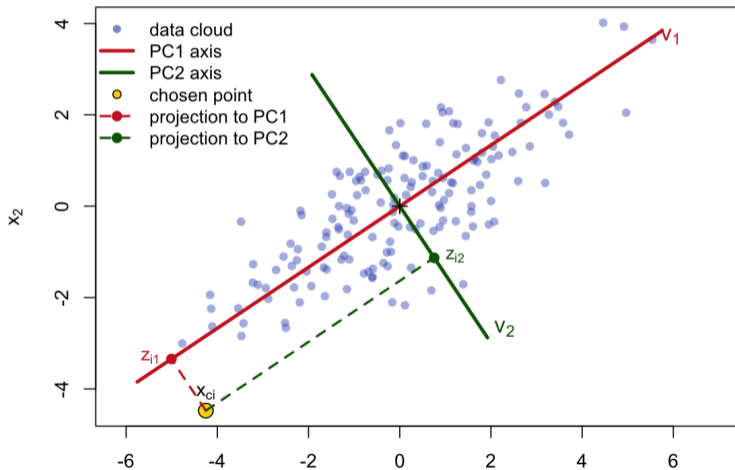
$$\mathbf{v}^\top \mathbf{S} \mathbf{v} = \sum_{j=2}^p \lambda_j c_j^2 \leq \lambda_2.$$

- Equality holds at $\mathbf{v} = \mathbf{v}_2$.

Therefore PC2 is an eigenvector for λ_2 (and similarly for later PCs).

Geometry of PC1 and PC2 axes

Geometry of PC1/PC2 axes and projections



Total variance and explained variance

For centered data $\mathbf{x}_{c,1}, \dots, \mathbf{x}_{c,p}$:

$$\text{Total sample variance} = \sum_{j=1}^p s_{jj} = \text{tr}(\mathbf{S}).$$

- Because diagonal entries s_{jj} are sample variances of each variable.

Using spectral decomposition $\mathbf{S} = \mathbf{V}\mathbf{\Lambda}\mathbf{V}^\top$:

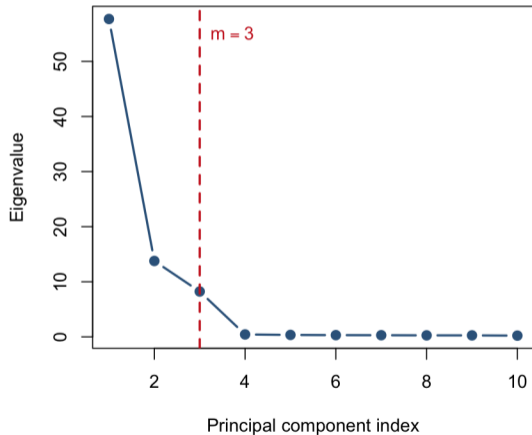
$$\text{tr}(\mathbf{S}) = \text{tr}(\mathbf{\Lambda}) = \sum_{j=1}^p \lambda_j.$$

So variance explained by the first m PCs is

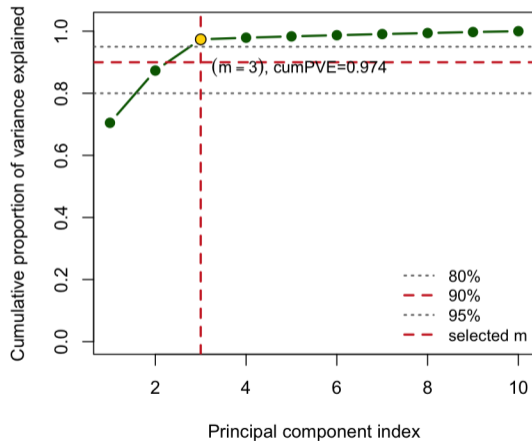
$$\text{cumPVE}(m) = \frac{\sum_{j=1}^m \lambda_j}{\sum_{j=1}^p \lambda_j}.$$

How many PC to select? scree plot + cumulative PVE

Scree plot



Cumulative PVE



Review of SVD

For centered data matrix $\mathbf{X}_c \in \mathbb{R}^{n \times p}$ with rank $r \leq \min(n, p)$:

$$\mathbf{X}_c = \mathbf{U}\mathbf{D}\mathbf{V}^\top,$$

where (thin SVD form)

$$\mathbf{U} \in \mathbb{R}^{n \times r}, \quad \mathbf{D} \in \mathbb{R}^{r \times r}, \quad \mathbf{V} \in \mathbb{R}^{p \times r}.$$

- $\mathbf{U}^\top \mathbf{U} = \mathbf{I}_r$, $\mathbf{V}^\top \mathbf{V} = \mathbf{I}_r$ (orthonormal columns).
- $\mathbf{D} = \text{diag}(d_1, \dots, d_r)$ with $d_1 \geq \dots \geq d_r > 0$ (singular values).
- Equivalent full SVD uses $\mathbf{U} \in \mathbb{R}^{n \times n}$, $\mathbf{D} \in \mathbb{R}^{n \times p}$, $\mathbf{V} \in \mathbb{R}^{p \times p}$.

Connection between SVD and PCA

Using $\mathbf{X}_c = UDV^\top$,

$$\mathbf{S} = \frac{1}{n-1} \mathbf{X}_c^\top \mathbf{X}_c = \frac{1}{n-1} VD^2V^\top.$$

- So columns of V are eigenvectors of \mathbf{S} (PC directions / loadings).
- Eigenvalues of \mathbf{S} are

$$\lambda_j = \frac{d_j^2}{n-1}, \quad j = 1, \dots, r.$$

- Scores are

$$\mathbf{Z} = \mathbf{X}_c V = UD,$$

score variance along PC j is λ_j .

- **Solving PCA using SVD:** SVD is numerically stable / faster computation when $n \ll p$.
- PCA naturally connects to low-rank approximations of the centered data matrix

Covariance PCA vs correlation PCA

- **Covariance PCA:** run PCA on \mathbf{X}_C .
- **Correlation PCA:** standardize each column to have variance 1 first, then PCA.

When to use which?

- Similar units/scale (e.g., yield changes in bps): covariance PCA often preferred.
- Mixed units/scales (returns, volume, macro levels): correlation PCA usually safer.

Reminder: choice changes the meaning of “variance explained”.

Dataset: daily changes in zero-coupon treasury yields at maturities

$\{1Y, 2Y, 3Y, 5Y, 7Y, 10Y, 20Y, 30Y\}$.

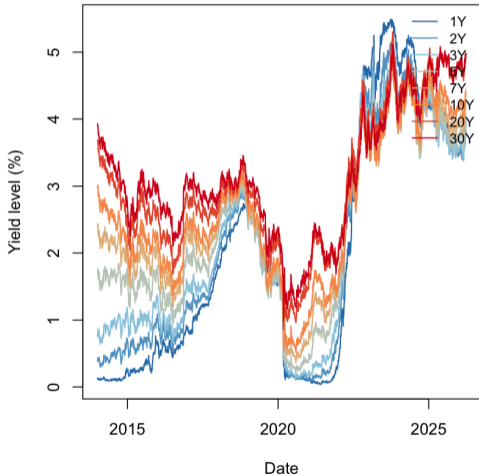
- FRED maturity series from 2014-2026

Modeling goal:

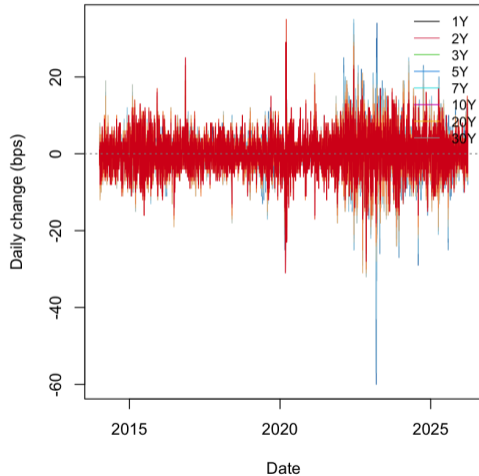
- Understand day-to-day curve changes for risk management.
- Use PCA to explain yield-curve movement.

Financial data example: treasury yield changes

Treasury yields by maturity (levels)



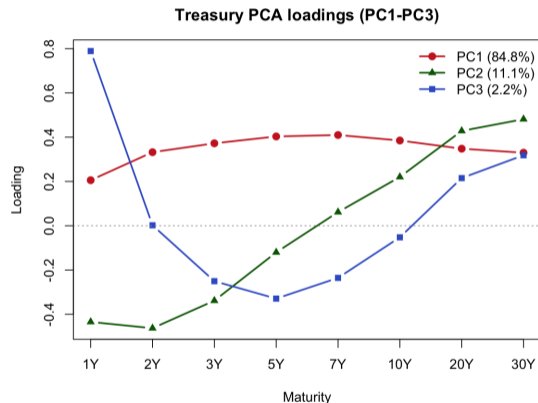
Daily yield changes by maturity



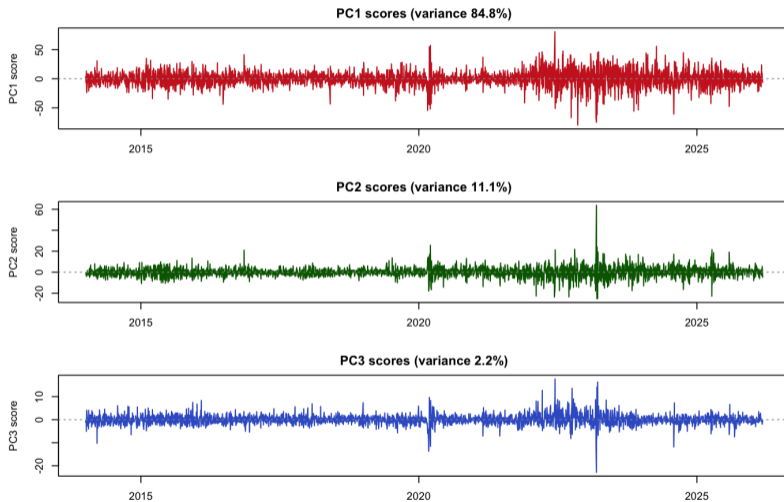
Interpretation of three PCs:

- PC1 (Level): nearly same-sign loadings across maturities.
- PC2 (Slope): short-end and long-end opposite signs.
- PC3 (Curvature): middle maturities opposite to ends.

These are called yield curve risk factors.



Financial data example: PCA scores



R code template (sample PCA on centered data X_c)

```
# X: n x p raw matrix/data.frame (rows = days, cols = maturities)
X_c <- scale(X, center = TRUE, scale = FALSE)

# Sample PCA via SVD (prcomp internally uses SVD)
pca <- prcomp(X_c, center = FALSE, scale. = FALSE)

eig <- pca$sdev^2
pve <- eig / sum(eig)
V   <- pca$rotation      # loading vectors v1, v2, ...
Z   <- pca$x             # scores = X_c %*% V
```

Python code template (sample PCA with consistent notation)

```
import numpy as np
from sklearn.decomposition import PCA

# X: n x p raw array
X_c = X - X.mean(axis=0, keepdims=True)

pca = PCA().fit(X_c)
eig = pca.explained_variance_
pve = pca.explained_variance_ratio_
V    = pca.components_.T    # columns are loading vectors v1, v2, ...
Z    = X_c @ V              # scores
```

- Population PCA (on Σ) v.s. sample PCA (on \mathbf{S})
- PCA directions maximize projected variance under unit-norm and orthogonality constraints
- Lagrangian + FOC lead naturally to eigenvectors/eigenvalues.
- Solving PCA using SVD
- In finance, yield-curve PCA connects directly to level/slope/curvature factors

- Johnson & Wichern (6th Edition), Chapter 8.1-8.4.